

Stochastic Calculus For Finance - volume 2
Section 4.6 Multivariable Stochastic Calculus
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Definition

4.6.1 Multiple Brownian Motions

Definition 4.6.1. A d -dimensional Brownian motion is a process

$$W(t) = (W_1(t), \dots, W_d(t))$$

with the following properties.

- (i) Each $W_i(t)$ is a one-dimensional Brownian motion.
- (ii) If $i \neq j$, then the processes $W_i(t)$ and $W_j(t)$ are independent.

Associated with a d -dimensional Brownian motion, we have a filtration $\mathcal{F}(t)$, $t \geq 0$, such that the following holds.

- (iii) **(Information accumulates)** For $0 \leq s < t$, every set in $\mathcal{F}(s)$ is also in $\mathcal{F}(t)$.
- (iv) **(Adaptivity)** For each $t \geq 0$, the random vector $W(t)$ is $\mathcal{F}(t)$ -measurable.
- (v) **(Independence of future increments)** For $0 \leq t < u$, the vector of increments $W(u) - W(t)$ is independent of $\mathcal{F}(t)$.

Notes

$$dW_i(t) dW_i(t) = dt. \quad dW_i(t) dW_j(t) = 0, \quad i \neq j.$$

proof:

Let $\Pi = \{t_0, \dots, t_n\}$ be a partition of $[0, T]$. For $i \neq j$, define the *sampled cross variation* of W_i and W_j on $[0, T]$ to be

$$C_\Pi = \sum_{k=0}^{n-1} [W_i(t_{k+1}) - W_i(t_k)] [W_j(t_{k+1}) - W_j(t_k)].$$

The increments appearing on the right-hand side of the equation above are all independent of one another and all have mean zero. Therefore, $\mathbb{E}C_\Pi = 0$.

proof(continue):

$$C_{\Pi} = \sum_{k=0}^{n-1} [W_i(t_{k+1}) - W_i(t_k)] [W_j(t_{k+1}) - W_j(t_k)].$$

We compute $\text{Var}(C_{\Pi})$. Note first that

$$\begin{aligned} C_{\Pi}^2 &= \sum_{k=0}^{n-1} [W_i(t_{k+1}) - W_i(t_k)]^2 [W_j(t_{k+1}) - W_j(t_k)]^2 \\ &\quad + 2 \sum_{\ell < k}^{n-1} [W_i(t_{\ell+1}) - W_i(t_{\ell})] [W_j(t_{\ell+1}) - W_j(t_{\ell})] \\ &\quad \cdot [W_i(t_{k+1}) - W_i(t_k)] [W_j(t_{k+1}) - W_j(t_k)]. \end{aligned}$$

All the increments appearing in the sum of cross-terms are independent of one another and all have mean zero. Therefore,

$$\text{Var}(C_{\Pi}) = \mathbb{E} C_{\Pi}^2 = \mathbb{E} \sum_{k=0}^{n-1} [W_i(t_{k+1}) - W_i(t_k)]^2 [W_j(t_{k+1}) - W_j(t_k)]^2.$$

But $[W_i(t_{k+1}) - W_i(t_k)]^2$ and $[W_j(t_{k+1}) - W_j(t_k)]^2$ are independent of one another, and each has expectation $(t_{k+1} - t_k)$. It follows that

$$\text{Var}(C_{\Pi}) = \sum_{k=0}^{n-1} (t_{k+1} - t_k)^2 \leq \|\Pi\| \cdot \sum_{k=0}^{n-1} (t_{k+1} - t_k) = \|\Pi\| \cdot T.$$

As $\|\Pi\| \rightarrow 0$, we have $\text{Var}(C_{\Pi}) \rightarrow 0$, so C_{Π} converges to the constant $\mathbb{E} C_{\Pi} = 0$.

Ito-Doeblin formula for Multiple processes(for 2-dimentional)

Let $X(t)$ and $Y(t)$ be Itô processes, which means they are processes of the form

$$X(t) = X(0) + \int_0^t \Theta_1(u) du + \int_0^t \sigma_{11}(u) dW_1(u) + \int_0^t \sigma_{12}(u) dW_2(u),$$

$$Y(t) = Y(0) + \int_0^t \Theta_2(u) du + \int_0^t \sigma_{21}(u) dW_1(u) + \int_0^t \sigma_{22}(u) dW_2(u).$$

The integrands $\Theta_i(u)$ and $\sigma_{ij}(u)$ are assumed to be adapted processes. In differential notation, we write

$$dX(t) = \Theta_1(t) dt + \sigma_{11}(t) dW_1(t) + \sigma_{12}(t) dW_2(t), \quad (4.6.1)$$

$$dY(t) = \Theta_2(t) dt + \sigma_{21}(t) dW_1(t) + \sigma_{22}(t) dW_2(t). \quad (4.6.2)$$

The Itô integral $\int_0^t \sigma_{11}(u) dW_1(u)$ accumulates quadratic variation at rate $\sigma_{11}^2(t)$ per unit time, and the Itô integral $\int_0^t \sigma_{12}(u) dW_2(u)$ accumulates quadratic variation at rate $\sigma_{12}^2(t)$ per unit time. Because both of these integrals appear in $X(t)$, the process $X(t)$ accumulates quadratic variation at rate $\sigma_{11}^2(t) + \sigma_{12}^2(t)$ per unit time:

$$[X, X](t) = \int_0^t (\sigma_{11}^2(u) + \sigma_{12}^2(u)) du.$$

We may write this equation in differential form as

$$dX(t) dX(t) = (\sigma_{11}^2(t) + \sigma_{12}^2(t)) dt. \quad (4.6.3)$$

proof:

$$\begin{aligned} dX(t) dX(t) &= \left(\theta_1(t) dt + \sigma_{11}(t) dW_1(t) + \sigma_{22}(t) dW_2(t) \right)^2 \\ &= \theta_1^2(t) dt^2 + \sigma_{11}^2(t) dW_1^2(t) + \sigma_{22}^2(t) dW_2^2(t) + 2\theta_1(t) dt \sigma_{11}(t) dW_1(t) \\ &\quad + 2\theta_1(t) dt \sigma_{22}(t) dW_2(t) + 2\sigma_{11}(t) \sigma_{22}(t) dW_1(t) dW_2(t) \\ &= \left[\sigma_{11}^2(t) + \sigma_{22}^2(t) \right] dt. \end{aligned}$$

Continue

In a similar way, we may derive the differential formulas

$$dY(t) dY(t) = (\sigma_{21}^2(t) + \sigma_{22}^2(t)) dt, \quad (4.6.4)$$

$$dX(t) dY(t) = (\sigma_{11}(t)\sigma_{21}(t) + \sigma_{12}(t)\sigma_{22}(t)) dt. \quad (4.6.5)$$

Equation (4.6.5) says that, for every $T \geq 0$,

$$[X, Y](T) = \int_0^T (\sigma_{11}(t)\sigma_{21}(t) + \sigma_{12}(t)\sigma_{22}(t)) dt. \quad (4.6.6)$$

Theorem 4.6.2. (Two-Dimensional Ito-Doeblin formula)

Theorem 4.6.2 (Two-dimensional Itô-Doeblin formula). *Let $f(t, x, y)$ be a function whose partial derivatives $f_t, f_x, f_y, f_{xx}, f_{xy}, f_{yx}$, and f_{yy} are defined and are continuous. Let $X(t)$ and $Y(t)$ be Itô processes as discussed above. The two-dimensional Itô-Doeblin formula in differential form is*

$$\begin{aligned} df(t, X(t), Y(t)) &= f_t(t, X(t), Y(t)) dt + f_x(t, X(t), Y(t)) dX(t) + f_y(t, X(t), Y(t)) dY(t) \\ &\quad + \frac{1}{2} f_{xx}(t, X(t), Y(t)) dX(t) dX(t) + f_{xy}(t, X(t), Y(t)) dX(t) dY(t) \\ &\quad + \frac{1}{2} f_{yy}(t, X(t), Y(t)) dY(t) dY(t). \end{aligned} \tag{4.6.8}$$

Before discussing formula (4.6.8), we rewrite it, leaving out t wherever possible, to obtain the same formula in the more compact notation

$$\begin{aligned} df(t, X, Y) &= f_t dt + f_x dX + f_y dY \\ &\quad + \frac{1}{2} f_{xx} dX dX + f_{xy} dX dY + \frac{1}{2} f_{yy} dY dY. \end{aligned} \tag{4.6.9}$$

proof:

$$\begin{aligned} df(t, X(t), Y(t)) &= f_t(t, X(t), Y(t)) dt + f_x(t, X(t), Y(t)) dX_t + f_y(t, X(t), Y(t)) dY_t \\ &+ \frac{1}{2} f_{tt}(t, X(t), Y(t)) dt^2 + \frac{1}{2} f_{xx}(t, X(t), Y(t)) dX_t^2 + \frac{1}{2} f_{yy}(t, X(t), Y(t)) dY_t^2 \\ &+ \frac{1}{2} (f_{ty} + f_{yt})(t, X(t), Y(t)) dt dY_t + \frac{1}{2} (f_{tx} + f_{xt})(t, X(t), Y(t)) dt dX_t + \frac{1}{2} (f_{xy} + f_{yx})(t, X(t), Y(t)) dX_t dY_t \end{aligned}$$

Since $dt^2 = 0$, $dt dY_t = dt dX_t = 0$.

$$\Rightarrow df(t, X(t), Y(t)) = f_t dt + f_x dX + f_y dY + \frac{1}{2} f_{xx} dX^2 + \frac{1}{2} f_{yy} dY^2 + f_{xy} dX dY$$

Note: Since f_{xy} , f_{yx} are both continuous and defined, so we have

$$\begin{aligned} f_{xy} &= \lim_{k \rightarrow 0} \frac{f_y(x+k, y) - f_y(x, y)}{k} = \lim_{k \rightarrow 0} \lim_{h \rightarrow 0} \frac{(f(x+k, y+h) - f(x+k, y)) - (f(x, y+h) - f(x, y))}{kh} \\ &= \lim_{h \rightarrow 0} \lim_{k \rightarrow 0} \frac{(f(x+k, y+h) - f(x, y+h)) - (f(x+k, y) - f(x, y))}{kh} = \lim_{h \rightarrow 0} \frac{f_x(x, y+h) - f_x(x, y)}{h} \\ &= f_{yx}. \end{aligned}$$

Corollary 4.6.3. (Ito product rule)

$$df(t, X, Y) = f_t dt + f_x dX + f_y dY + \frac{1}{2}f_{xx} dX dX + f_{xy} dX dY + \frac{1}{2}f_{yy} dY dY. \quad (4.6.9)$$

Corollary 4.6.3 (Itô product rule). *Let $X(t)$ and $Y(t)$ be Itô processes. Then*

$$d(X(t)Y(t)) = X(t) dY(t) + Y(t) dX(t) + dX(t) dY(t).$$

PROOF: In (4.6.9), take $f(t, x, y) = xy$, so that $f_t = 0$, $f_x = y$, $f_y = x$, $f_{xx} = 0$, $f_{xy} = 1$, and $f_{yy} = 0$. \square

Theorem 4.6.4. (Levy's One-Dimensional Characterization Theorem)

Theorem 4.6.4 (Lévy, one dimension). *Let $M(t)$, $t \geq 0$, be a martingale relative to a filtration $\mathcal{F}(t)$, $t \geq 0$. Assume that $M(0) = 0$, $M(t)$ has continuous paths, and $[M, M](t) = t$ for all $t \geq 0$. Then $M(t)$ is a Brownian motion.*

proof(idea):

- Lemma
- Euler formula: $\exp(ix) = \cos(x) + i\sin(x)$, where i is imaginary
- $M(t) - M(s) \sim N(0, t-s)$
- $M(t_1) - M(t_2), M(t_2) - M(t_3), M(t_3) - M(t_4), \dots$ are independent

proof:

Lemma Let $M(t)$ be as in Thm 4.6.4 and let $f \in C(\mathbb{R}^2)$ such that f, f', f'' are bounded

$\forall 0 \leq s \leq t$, we have $E(f(M_t) | \mathcal{F}_s) = f(M_s) + \frac{1}{2} \int_s^t E(f''(M_u) | \mathcal{F}_s) du$

pf Let $\pi = \{s = t_0 < t_1 < \dots < t_n = t\}$ be a partition of $[s, t]$. By Taylor's formula:

$f(X_t) = f(X_s) + \sum_{k=1}^n (f(X_{t_k}) - f(X_{t_{k-1}}))$, and applying Taylor's formula on $f(X_{t_k}) - f(X_{t_{k-1}})$

$$f(M_t) = f(M_s) + \sum_{k=1}^n f'(M_{t_{k-1}})(M_{t_k} - M_{t_{k-1}}) + \frac{1}{2} \sum_{k=1}^n f''(M_{t_{k-1}})(M_{t_k} - M_{t_{k-1}})^2 + R_{\pi} \quad (*)$$

We note that as $\|\pi\| \rightarrow 0$, $R_{\pi} \rightarrow 0$ since $[M, M](t) = t$, so $(M_{t_k} - M_{t_{k-1}})^2$ for $k \geq 3$

$$= (M_{t_k} - M_{t_{k-1}})^2 \cdot (M_{t_k} - M_{t_{k-1}})^{k-2} = (t_k - t_{k-1})(M_{t_k} - M_{t_{k-1}})^{k-2} \leq \|\pi\| \cdot (M_{t_k} - M_{t_{k-1}})^{k-2} = 0.$$

Since M_t has continuous path, $M_{t_k} - M_{t_{k-1}} \leq \infty$ for bounded interval.

By Law of total expectation (For $\mathcal{F}_s \subseteq \mathcal{F}_t \subseteq \mathcal{F}$, $E(E(X | \mathcal{F}_t) | \mathcal{F}_s) = E(X | \mathcal{F}_s)$)

Take conditional expectation on \mathcal{F}_s to $(*)$ we have:

$$E(f(M_t) | \mathcal{F}_s) = f(M_s) + \sum_{k=1}^n E(E(f'(M_{t_{k-1}})(M_{t_k} - M_{t_{k-1}})) | \mathcal{F}_{t_{k-1}} | \mathcal{F}_s) + \frac{1}{2} \sum_{k=1}^n E(E(f''(M_{t_{k-1}}))$$

$$\cdot (M_{t_k} - M_{t_{k-1}})^2 | \mathcal{F}_{t_{k-1}} | \mathcal{F}_s) = f(M_s) + \frac{1}{2} \sum_{k=1}^n E(f''(M_{t_{k-1}}) | \mathcal{F}_s) (t_k - t_{k-1}).$$

since M_t has continuous path, $E((M_{t_k} - M_{t_{k-1}}) | \mathcal{F}_{t_{k-1}}) = 0$, $E(E(f''(M_{t_{k-1}})(M_{t_k} - M_{t_{k-1}})^2 | \mathcal{F}_{t_{k-1}}) | \mathcal{F}_s)$

$$= E(f''(M_{t_{k-1}})(t_k - t_{k-1}) | \mathcal{F}_s) = E(f''(M_{t_{k-1}}) | \mathcal{F}_s) (t_k - t_{k-1}).$$

$$\text{As } \|\pi\| \rightarrow 0, f(M_s) + \frac{1}{2} \sum_{k=1}^n E(f''(M_{t_{k-1}}) | \mathcal{F}_s) (t_k - t_{k-1}) \rightarrow f(M_s) + \frac{1}{2} \int_s^t E(f''(M_u) | \mathcal{F}_s) du.$$

proof(continue):

pt) We would like to show $\forall 0 \leq s \leq t, E(\exp(iu(M_t - M_s)) | \mathcal{F}_s) = \exp(-\frac{1}{2}u^2(t-s)), \forall u \in \mathbb{R}$.
 let $f(x) = \exp(iux)$. $\forall s \leq t, E(\exp(iuM_t) | \mathcal{F}_s) = \exp(iuM_s) - \frac{1}{2}u^2 \int_s^t E(\exp(iuM_u) | \mathcal{F}_s) du$, since $f'(x) = u(\sin ux - i \cos ux)$, $f''(x) = u^2(-\cos ux - i \sin ux) = -u^2 \exp(iux)$.
 Also (*) $\Rightarrow E(\exp(iu(M_t - M_s)) | \mathcal{F}_s) = 1 - \frac{1}{2}u^2 \int_s^t E(\exp(iu(M_u - M_s)) | \mathcal{F}_s) du$.

$$\Rightarrow \frac{d}{dt} E(\exp(iu(M_t - M_s)) | \mathcal{F}_s) = -\frac{1}{2}u^2 E(\exp(iu(M_t - M_s)) | \mathcal{F}_s)$$

let $g(t) = E(\exp(iu(M_t - M_s)) | \mathcal{F}_s)$, we have $g'(t) = -\frac{1}{2}u^2 g(t)$, $g(s) = 1$

Solving this ODE gives $E(\exp(iu(M_t - M_s)) | \mathcal{F}_s) = g(t) = \exp(-\frac{1}{2}u^2(t-s))$.

If this can be shown, then we have $M_t - M_s \sim N(0, t-s)$, since the characteristic function of $M_t - M_s = E(\exp(iu(M_t - M_s))) = E(E(\exp(iu(M_t - M_s)) | \mathcal{F}_s)) = E(\exp(-\frac{1}{2}u^2(t-s))) = \exp(-\frac{1}{2}u^2(t-s))$.

Again by Law of Iterated Expectations and given the partition $0 = t_0 < t_1 < \dots < t_n = t < \infty$.

$$E\left(\prod_{i=1}^n \exp(iu(M_{t_i} - M_{t_{i-1}})) | \mathcal{F}_0\right) = E\left(E\left(\prod_{i=1}^n \exp(iu(M_{t_i} - M_{t_{i-1}})) | \mathcal{F}_{t_{i-1}}\right) | \mathcal{F}_0\right)$$

$$= (t_n - t_{n-1}) \cdot E\left(E\left(\prod_{i=1}^{n-1} \exp(iu(M_{t_i} - M_{t_{i-1}})) | \mathcal{F}_{t_{i-1}}\right) | \mathcal{F}_0\right) = \dots = \prod_{i=1}^n (t_i - t_{i-1})$$

$$= \prod_{i=1}^n E(\exp(iu(M_{t_i} - M_{t_{i-1}}))) \Rightarrow \phi_Y = \prod_{i=1}^n \phi_{M_{t_i} - M_{t_{i-1}}}, Y = (M_{t_1} - M_{t_0}, M_{t_2} - M_{t_1}, \dots, M_{t_n} - M_{t_{n-1}})$$

$\Rightarrow M_{t_1} - M_{t_0}, M_{t_2} - M_{t_1}, \dots, M_{t_n} - M_{t_{n-1}}$ are all independent.